

Lecture 9

November 3, 2010

Bounds on Estimates:

Cramer - Rao Lower Bound (CRLB).

Assume that $\hat{\theta}(y)$ is an unbiased estimate of θ .

Then:

$$E[\hat{\theta}(y) - \theta] = \int_{\Gamma} [\hat{\theta}(y) - \theta] p(y|\theta) dy = 0$$

Taking derivative with respect to θ , we get:

$$\begin{aligned} \frac{d}{d\theta} \int_{\Gamma} [\hat{\theta}(y) - \theta] p(y|\theta) dy &= \int_{\Gamma} \frac{\partial}{\partial \theta} \{ [\hat{\theta}(y) - \theta] p(y|\theta) \} dy \\ &= - \int_{\Gamma} p(y|\theta) dy + \int_{\Gamma} \frac{\partial p(y|\theta)}{\partial \theta} [\hat{\theta}(y) - \theta] dy = 0 \end{aligned}$$

or

$$\int_{\Gamma} \frac{\partial}{\partial \theta} p(y|\theta) [\hat{\theta}(y) - \theta] dy = 1$$

We can write this as:

$$\int_{\Gamma} \frac{\partial \log p(y|\theta)}{\partial \theta} p(y|\theta) [\hat{\theta}(y) - \theta] dy = 1$$

where we have used the fact that:

$$\frac{d \log f(x)}{dx} = \frac{\frac{df(x)}{dx}}{f(x)} \Rightarrow f(x) = f(x) \frac{d \log f(x)}{dx}$$

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This further can be written as:

$$\int_{\Gamma} \left[\frac{\partial \log p(y|\theta)}{\partial \theta} \sqrt{p(y|\theta)} \right] \left[\sqrt{p(y|\theta)} (\hat{\theta}(y) - \theta) \right] dy = 1$$

Using Schwarz's inequality, i.e.,

$$\left[\int g(x) h(x) dx \right]^2 \leq \int g^2(x) dx \int h^2(x) dx$$

with equality if $g(x) = k h(x)$,

we get:

$$\int_{\Gamma} \left[\frac{\partial \log p(y|\theta)}{\partial \theta} \right]^2 p(y|\theta) dy \int_{\Gamma} (\hat{\theta}(y) - \theta)^2 p(y|\theta) dy \geq 1$$

with equality if and only if: $\text{Var}[\hat{\theta}(y) - \theta]$

$$\frac{\partial \log p(y|\theta)}{\partial \theta} = [\hat{\theta} - \theta] k(\theta)$$

where $k(\theta)$ is not a function of y .

Finally, we have

$$\text{Var}[\hat{\theta} - \theta] = E \{ (\hat{\theta}(y) - \theta)^2 | \theta \} \geq \frac{1}{E \left\{ \left[\frac{\partial \log p(y|\theta)}{\partial \theta} \right]^2 \right\} | \theta}$$

This is called the Cramer-Rao Lower Bound.

[Note: We have denoted natural logarithm by \log and not \ln .]

Alternative form of CRLB:

We have

$$\int p(y|\theta) dy = 1$$

Taking derivative:

$$\int \frac{\partial}{\partial \theta} p(y|\theta) dy = \int \frac{\partial \log p(y|\theta)}{\partial \theta} p(y|\theta) d\theta = 0$$

Taking derivative again:

$$0 = \int \frac{\partial^2 \log p(y|\theta)}{\partial \theta^2} p(y|\theta) d\theta + \int \left[\frac{\partial \log p(y|\theta)}{\partial \theta} \right]^2 p(y|\theta) d\theta$$

or

$$E \left[\frac{\partial^2 \log p(y|\theta)}{\partial \theta^2} \middle| \theta \right] = - E \left[\left(\frac{\partial \log p(y|\theta)}{\partial \theta} \right)^2 \middle| \theta \right]$$

So:

$$E[(\hat{\theta}(y) - \theta)^2 | \theta] \geq \frac{-1}{E \left[\frac{\partial^2 \log p(y|\theta)}{\partial \theta^2} \middle| \theta \right]}$$

An estimate that achieves CRLB is called an efficient estimate.

For an efficient estimate:

$$\frac{\partial \log p(y|\theta)}{\partial \theta} = [\hat{\theta}(y) - \theta] k(\theta)$$

Assume that an efficient estimate exists, i.e., we have:

$$\frac{\partial}{\partial \theta} \log p(y|\theta) = [\hat{\theta}(y) - \theta] k(\theta)$$

Substitute θ by $\hat{\theta}_{ML}$, i.e.,

$$\left. \frac{\partial}{\partial \theta} \log p(y|\theta) \right|_{\theta = \hat{\theta}_{ML}} = [\hat{\theta}(y) - \hat{\theta}_{ML}] k(\theta)$$

where $\hat{\theta}_{ML}$ is the maximum-likelihood estimate.

But,

$$\left. \frac{\partial}{\partial \theta} \log p(y|\theta) \right|_{\theta = \hat{\theta}_{ML}} = 0$$

So,

$$[\hat{\theta}(y) - \hat{\theta}_{ML}] k(\theta) = 0$$

or
$$\hat{\theta}(y) = \hat{\theta}_{ML}$$

This means, if there is an efficient estimate then it is ML estimate and can be found by solving

$$\left. \frac{\partial}{\partial \theta} \log p(y|\theta) \right|_{\theta = \hat{\theta}_{ML}} = 0.$$

But if an efficient estimate does not exist, we cannot know how ^{good} the ML estimate is.

Example: Take the example that we considered in the previous lecture, i.e.,

$$p(\underline{y} | \theta) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi}\sigma} \exp\left[-\frac{(y_i - \theta)^2}{2\sigma^2}\right]$$

Then

$$\frac{\partial \log p(\underline{y} | \theta)}{\partial \theta} = \frac{1}{\sigma^2} \left(\sum_{i=1}^n y_i - N\theta \right)$$

So,

$$\hat{\theta}_{ML} = \frac{1}{n} \sum_{i=1}^n y_i$$

and

$$E[\hat{\theta}_{ML}] = \frac{1}{n} \sum_{i=1}^n E\{y_i\} = \frac{1}{n} \sum_{i=1}^n \theta = \theta$$

So $\hat{\theta}_{ML}$ is unbiased.

Since the condition of Cramer-Rao bound equality is satisfied, then the variance of the estimate is given by CRLB.

$$\frac{\partial^2 \log p(\underline{y} | \theta)}{\partial \theta^2} = -\frac{N}{\sigma^2}$$

So:

$$\text{Var}[\hat{\theta}_{ML} - \theta] = \frac{\sigma^2}{N}$$

Example: Now consider the estimation of a non random parameter of a signal of known form, e.g., the phase or frequency of a sinusoid.

$$y(t) = g(t, \theta) + N(t) \quad t_0 \leq t \leq t_f$$

Let's assume that we have samples of the received waveform $y(t)$ at points $t_k = kT, k=1, 2, \dots$

Then

$$y_k = g_k + N_k$$

where $g_k = g(t_k, \theta)$

It is clear that

$$E\{y_k\} = E\{g_k\}$$

and

$$\text{Var}\{y_k | \theta\} = \text{Var}(N_k) = \sigma^2$$

Let vector \underline{y}_n , be equal to $[y_1, \dots, y_n]^T$

Then

$$\begin{aligned} p(\underline{y}_n | \theta) &= \prod_{i=1}^n p(y_i | \theta) \\ &= \prod_{i=1}^n \frac{1}{\sqrt{2\pi}\sigma} \exp\left[-\frac{(y_i - g_i)^2}{2\sigma^2}\right] \\ &= \left(\frac{1}{\sqrt{2\pi}\sigma}\right)^n \exp\left[-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - g_i)^2\right] \end{aligned}$$

We have

$$\frac{\partial \log P(\underline{y}_n | \theta)}{\partial \theta} = \frac{1}{\sigma^2} \sum_{i=1}^n (y_i - g_i) \frac{\partial g_i}{\partial \theta}$$

if we let $n \rightarrow \infty$, then the summation tends to an integral and the ML estimate will be the solution to:

$$\int_{t_0}^{t_f} [y(t) - g(t, \theta)] \frac{\partial g(t, \theta)}{\partial \theta} dt \Big|_{\theta = \hat{\theta}_{ML}} = 0$$

Example: Amplitude estimation:

$$y(t, \theta) = \theta y(t)$$

For this case

$$\int_{t_0}^{t_f} [y(t) - \theta g(t)] g(t) dt \Big|_{\theta = \hat{\theta}_{ML}} = 0$$

So

$$\hat{\theta}_{ML} = \frac{\int_{t_0}^{t_f} y(t) g(t) dt}{\int_{t_0}^{t_f} g^2(t) dt}$$

Normalizing $g(t)$, i.e., letting:

$$\int_{t_0}^{t_f} g^2(t) dt = 1$$

We get :

$$\hat{\theta}_{ML} = \int_{t_0}^{t_f} y(t) g(t) dt$$

which is the correlation receiver.

We have

$$E\{\hat{\theta}_{ML}\} = E\left\{ \int_{t_0}^{t_f} [\theta g(t) + N(t)] g(t) dt \right\} = \theta$$

So, the estimate is unbiased.

$$\begin{aligned} \frac{\partial \log p(y|\theta)}{\partial \theta} &= \frac{2}{N_0} \int_{t_0}^{t_f} [y(t) - \theta g(t)] g(t) dt \\ &= \frac{2}{N_0} [\hat{\theta}_{ML} - \theta] \end{aligned}$$

So $\hat{\theta}_{ML}$ is efficient, i.e., it achieves CRLB.

Here $N_0 = \frac{\sigma^2}{W}$ where W is the bandwidth of the signal.

Example: Phase estimation

$$g(t, \theta) = A \sin(\omega_0 t + \theta)$$

$$\int_{t_0}^{t_f} y(t) \cos(\omega_0 t + \hat{\theta}_{ML}) dt = 0$$

$$\text{or } \hat{\theta}_{ML} = \tan^{-1} \frac{\int_{t_0}^{t_f} y(t) \cos \omega_0 t dt}{\int_{t_0}^{t_f} y(t) \sin \omega_0 t dt}$$

CRLB for phase estimation.

By finding $E\left\{\frac{\partial^2 p(y|\theta)}{\partial \theta^2}\right\}$ and substituting in the Cramer-Rao Bound formula, we get

$$\text{Var}(\hat{\theta} - \theta) \geq \left[\frac{2E_s}{N_0}\right]^{-1} \frac{1}{n}$$

where E_s is the signal energy $\frac{A^2}{2}T$ and N_0 is the one sided noise density. So, the estimation improves as the length of observation increases and/or SNR increases.

For frequency estimation:

$$\text{Var}(\hat{\Omega} - \Omega) \geq \left[\frac{2E_s}{N_0}\right]^{-1} \frac{12}{(n-1)^2 n}$$